

# Approximation of Expectation and Variance on $[a, b]$ Interval, with Probability Density Function in $L_p[a, b], 0 < p < 1$

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## Abstract :

In this article we use Taylor's formula to approximate expression in terms of expectation and variance simultaneously with probability density function in  $L_p, 0 < p < 1$ .

### 1.1. Introduction

If  $\mathbf{X}$  is a random variable, have probability density function

$f: [a, b] \rightarrow R$ . We know that the expectation of the random variable  $\mathbf{X}$  is

$$E(X) = \int_{\infty}^{\infty} tf(t)dt.$$

Therefore, the variance of the random variable  $X$  is

$$\sigma^2 = \int_a^b (t - E(X)f(t)dt)^2 = E(X^2) - (E(X))^2$$

In our article we use these inequalities to approximate (estimate) expectation and variance with measurable probability density functions, in the aid of Taylor's formula.

### 1.2. The Main Results

To prove our main theorem we need the following auxiliary Lemmas

**Lemma 1.2.1:**[2]

$$\int_a^b (b-t)(t-a)f(t)dt = |b-E(X)||E(X)-a| - \sigma^2(X), t \in [a, b]$$

**Lemma 1.2.2.**[1]

If  $p < q$ , then

$$\left( \sum_{i=1}^{\infty} |x_i|^q \right)^{\frac{1}{q}} \leq \left( \sum_{i=1}^{\infty} |x_i|^p \right)^{\frac{1}{p}}.$$

Now let us introduce our main Theorem.

**Theorem 1.2.3**

Let  $\mathbf{X}$  be a random variable defined on  $[a, b]$  with the probability density function  $f: [a, b] \rightarrow R$  belongs to  $L_p[a, b], p < 1$ . Then we have,

$$|b - E(X)||E(X) - a| - \sigma^2(X) \leq c(b - a)^{2+p-\frac{1}{p}} \|f\|_p,$$

$0 < p < 1$ , where  $c$  is a positive constant.

## Proof

Since

$$\begin{aligned} \int_a^b (\mathbf{b} - \mathbf{t})(\mathbf{t} - \mathbf{a})f(t)dt &\leq \int_a^b (\mathbf{b} - \mathbf{a})^2 |f(t)|dt \\ &= (\mathbf{b} - \mathbf{a})^2 \int_a^b |f(t)|dt, \end{aligned} \quad (1)$$

Now

Let  $\mathbf{t}_1 < \mathbf{t}_2 < \mathbf{t}_3 < \dots < \mathbf{t}_n$  be a partition for  $[\mathbf{a}, \mathbf{b}]$ ,  $\mathbf{a} = \mathbf{t}_0$ ,  $\Delta = \frac{\mathbf{b}-\mathbf{a}}{n}$ ,

$$\mathbf{t}_1 = \mathbf{a} + \frac{\mathbf{b}-\mathbf{a}}{n}, \mathbf{t}_2 = \mathbf{a} + \frac{2(\mathbf{b}-\mathbf{a})}{n}, \dots, \mathbf{t}_n = \mathbf{a} + \frac{i(\mathbf{b}-\mathbf{a})}{n}.$$

This implies,

$$\int_a^b f(t)dt \cong \sum_{i=1}^n f(t_i) \frac{\mathbf{b}-\mathbf{a}}{n}, \quad [54] \quad (2)$$

If  $p < 1$ , by using Lemma 1.2.2 and (2), we get

$$\int_a^b f(t)dt \leq \left( \int_a^b |f(t)|^{\frac{1}{p}} dt \right)^p \leq c \left( \sum_{i=1}^n |f(t_i)|^{\frac{1}{p}} \frac{\mathbf{b}-\mathbf{a}}{n} \right)^p$$

By using Holder inequality when  $q > 1, k > 1$  and  $\frac{1}{q} + \frac{1}{k} = 1$ , we get

$$\begin{aligned} \int_a^b f(t)dt &\leq c(\mathbf{b}-\mathbf{a})^p \left( \left( \sum_{i=1}^n |f(t_i)|^{\frac{q}{p}} \right)^{\frac{p}{q}} \left( \sum_{i=1}^n \left| \frac{1}{n} \right|^k \right)^{\frac{p}{k}} \right) \\ &\leq c(\mathbf{b}-\mathbf{a})^p \left( \left( \sum_{i=1}^n |f(t_i)|^{\frac{q}{p}} \right)^{\frac{p}{q}} \sum_{i=1}^n \frac{1}{n^k} \right), k > 1 \\ &\leq c(\mathbf{b}-\mathbf{a})^p \left( \left( \sum_{i=1}^n |f(t_i)|^p \right)^{\frac{1}{p}} \frac{1}{n^{k-1}} \right) \end{aligned}$$

Assume  $\frac{1}{p} = k - 1$ , then

$$\begin{aligned} \int_a^b f(t)dt &\leq \frac{c(\mathbf{b}-\mathbf{a})^p}{(\mathbf{b}-\mathbf{a})^{\frac{1}{p}}} \left( \left( \sum_{i=1}^n |f(t_i)|^p \frac{(\mathbf{b}-\mathbf{a})}{n} \right)^{\frac{1}{p}} \right) \\ &\leq \frac{c(\mathbf{b}-\mathbf{a})^p}{(\mathbf{b}-\mathbf{a})^{\frac{1}{p}}} \left( \left( \int_a^b |f(t_i)|^p dt \right)^{\frac{1}{p}} \right) \\ &\leq \frac{c(\mathbf{b}-\mathbf{a})^p}{(\mathbf{b}-\mathbf{a})^{\frac{1}{p}}} \|f\|_p. \end{aligned} \quad (3)$$

Thus

$$\int_a^b (\mathbf{b} - \mathbf{t})(\mathbf{t} - \mathbf{a})f(t)dt \leq c(\mathbf{b}-\mathbf{a})^{2+p-\frac{1}{p}} \|f\|_p$$

Then by repeating of using of (2), we get,

$$\int_a^b (\mathbf{b} - \mathbf{t})(\mathbf{t} - \mathbf{a})f(t)dt \leq (\mathbf{b}-\mathbf{a})^{2+p-\frac{1}{p}} \|f\|_p.$$

Then by using Lemma 1.2.1 , we get,

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Vol.7 No.3 (March, 2022)

$$|\mathbf{b} - E(X)| |E(X) - \mathbf{a}| - \sigma^2(X) \leq (\mathbf{b} - \mathbf{a})^{2+p-\frac{1}{p}} \|f\|_p,$$

**0 < p < 1**

■

### Theorem 1.2.4

Let  $\mathbf{X}$  be a random variable defined in  $[\mathbf{a}, \mathbf{b}]$  with the probability density function  $f: [\mathbf{a}, \mathbf{b}] \rightarrow \mathbf{R}$  belongs to  $L_p[\mathbf{a}, \mathbf{b}]$ ,  $p < 1$ . Then we have,

$$\left| |\mathbf{b} - E(\mathbf{X})| |E(\mathbf{X}) - \mathbf{a}| - \sigma^2(\mathbf{X}) - \frac{(\mathbf{b} - \mathbf{a})^3}{6} \right| \leq \frac{(\mathbf{b} - \mathbf{a})}{2} ((\mathbf{a} + \mathbf{x})^2 + (\mathbf{b} + \mathbf{x})^2 \|f(\mathbf{t})\|_p) + \frac{(\mathbf{b} - \mathbf{a})^3}{6}$$

**Proof**

Recall pre-Gruss inequality when  $0 < p < 1$

$$\begin{aligned} & \left| \frac{1}{\mathbf{b} - \mathbf{a}} \int_a^b h(t)g(x, t)dt - \int_a^b h(t)dt \frac{1}{\mathbf{b} - \mathbf{a}} \int_a^b g(x, t)dt \right| \\ & \leq \frac{1}{2} \{(\mathbf{a} + \mathbf{x})^2 + (\mathbf{b} + \mathbf{x})^2 \|h(t)\|_p\} + \left| \int_a^b h(t)dt \frac{1}{\mathbf{b} - \mathbf{a}} \int_a^b g(x, t)dt \right| \quad (4) \end{aligned}$$

Put  $h(t) = f(t)$ ,  $g(x, t) = (\mathbf{b} - \mathbf{t})(\mathbf{t} - \mathbf{a})$ , in (4), we get.

$$\begin{aligned} & \left| \frac{1}{\mathbf{b} - \mathbf{a}} \int_a^b f(t)(\mathbf{b} - \mathbf{t})(\mathbf{t} - \mathbf{a})dt - \int_a^b f(t)dt \frac{1}{\mathbf{b} - \mathbf{a}} \int_a^b (\mathbf{b} - \mathbf{t})(\mathbf{t} - \mathbf{a})dt \right| \\ & \leq \frac{1}{2} ((\mathbf{a} + \mathbf{x})^2 + (\mathbf{b} + \mathbf{x})^2 \|f(t)\|_p) + \left| \int_a^b f(t)dt \frac{1}{\mathbf{b} - \mathbf{a}} \int_a^b (\mathbf{b} - \mathbf{t})(\mathbf{t} - \mathbf{a})dt \right|. \quad (5) \end{aligned}$$

Now let us compute  $\int_a^b (\mathbf{b} - \mathbf{t})(\mathbf{t} - \mathbf{a})dt$

$$\begin{aligned} & \int_a^b (\mathbf{b} - \mathbf{t})(\mathbf{t} - \mathbf{a})dt \\ & = \int_a^b (bt - ab - t^2 + at)dt = \frac{t^3}{3} - abt - \frac{t^3}{3} + \frac{at^2}{2} \Big|_a^b \\ & = \left( \frac{b^3}{2} - ab^2 - \frac{b^3}{3} + \frac{ab^2}{2} \right) - \left( \frac{a^2b}{2} - ba^2 - \frac{a^3}{3} + \frac{a^3}{2} \right) \\ & = \left( \frac{3b^3 - 6ab^2 - 2b^3 + 3ab^2}{6} \right) - \left( \frac{3a^2b - 6ba^2 - 2a^3 + 3a^3}{6} \right) \\ & = \frac{(b^3 - 3ab^2) - (a^3 - 3ba^2)}{6} = \frac{b^3 - 3ab^2 + 3ba^2 - a^3}{6} = \frac{(\mathbf{b} - \mathbf{a})^3}{6} \quad (6) \end{aligned}$$

Then using (6) to complete our estimate in (5). Also we have  $\int_a^b f(t)dt = 1$ , we obtain

$$\begin{aligned} & \left| \frac{1}{\mathbf{b} - \mathbf{a}} \int_a^b f(t)(\mathbf{b} - \mathbf{t})(\mathbf{t} - \mathbf{a})dt - \frac{(\mathbf{b} - \mathbf{a})^3}{6} \right| \\ & \leq \frac{1}{2} ((\mathbf{a} + \mathbf{x})^2 + (\mathbf{b} + \mathbf{x})^2 \|f(t)\|_p) + \frac{(\mathbf{b} - \mathbf{a})^2}{6}. \end{aligned}$$

Then,

$$\begin{aligned} & \left| \int_a^b f(t)(\mathbf{b} - \mathbf{t})(\mathbf{t} - \mathbf{a})dt - \frac{(\mathbf{b} - \mathbf{a})^3}{6} \right| \\ & \leq \frac{(\mathbf{b} - \mathbf{a})}{2} ((\mathbf{a} + \mathbf{x})^2 + (\mathbf{b} + \mathbf{x})^2 \|f(t)\|_p) + \frac{(\mathbf{b} - \mathbf{a})^3}{6}, \end{aligned}$$

Now by using Lemma 1.2.1 , we get,

$$\left| |\mathbf{b} - \mathbf{E}(X)| |\mathbf{E}(X) - \mathbf{a}| - \sigma^2(X) - \frac{(\mathbf{b} - \mathbf{a})^3}{6} \right| \leq \frac{(\mathbf{b} - \mathbf{a})}{2} ((\mathbf{a} + \mathbf{x})^2 + (\mathbf{b} + \mathbf{x})^2 \|f(t)\|_p) + \frac{(\mathbf{b} - \mathbf{a})^3}{6} \quad \blacksquare$$

#### Corollary 1.2.4

$$\left| |\mathbf{b} - \mathbf{E}(X)| |\mathbf{E}(X) - \mathbf{a}| - \sigma^2(X) - \frac{(\mathbf{b} - \mathbf{a})^3}{6} \right| \leq (\mathbf{b} + \mathbf{a})^3 (1 + \|f\|_p)$$

#### Proof:

Since,

By using Theorem 1.2.4, we get,

$$\left| |\mathbf{b} - \mathbf{E}(X)| |\mathbf{E}(X) - \mathbf{a}| - \sigma^2(X) - \frac{(\mathbf{b} - \mathbf{a})^3}{6} \right| \leq \frac{(\mathbf{b} - \mathbf{a})}{2} ((\mathbf{a} + \mathbf{x})^2 + (\mathbf{b} + \mathbf{x})^2 \|f(t)\|_p) + \frac{(\mathbf{b} - \mathbf{a})^3}{6} \leq (\mathbf{b} + \mathbf{a})^3 (1 + \|f\|_p) \quad \blacksquare$$

#### Theorem 1.2.5

Let  $\mathbf{X}$  be a random variable defined in  $[\mathbf{a}, \mathbf{b}]$  with the probability density function  $f: [\mathbf{a}, \mathbf{b}] \rightarrow \mathbf{R}$ . If  $f, \hat{f} \in L_p[\mathbf{a}, \mathbf{b}], p < 1$ , then we have,

$$|\mathbf{b} - \mathbf{E}(X)| |\mathbf{E}(X) - \mathbf{a}| - \sigma^2(X) \leq 2^{\frac{1}{p}+3} \mathbf{b}^{\frac{4}{p}} \|\hat{f}\|_p, 0 < p < 1$$

#### Proof:

Recall the "Pre-Chebychev inequality"

$$\begin{aligned} \left| \frac{1}{\mathbf{b} - \mathbf{a}} \int_a^b f(x) g(x) dx - \frac{1}{\mathbf{b} - \mathbf{a}} \int_a^b f(x) dx \frac{1}{\mathbf{b} - \mathbf{a}} \int_a^b g(x) dx \right| \leq \\ (\mathbf{b} - \mathbf{a}) \|\hat{f}\|_p \left| \frac{1}{(\mathbf{b} - \mathbf{a})} \int_a^b |g(x)|^2 dx - \left( \frac{1}{(\mathbf{b} - \mathbf{a})} \int_a^b |g(x)| dx \right)^2 \right|^{\frac{1}{p}} \end{aligned} \quad (7)$$

In (7) put  $g(x) = (\mathbf{t} - \mathbf{a})(\mathbf{b} - \mathbf{t})$ , we get

$$\begin{aligned} \left| \frac{1}{\mathbf{b} - \mathbf{a}} \int_a^b f(t)(\mathbf{t} - \mathbf{a})(\mathbf{b} - \mathbf{t}) dt - \frac{1}{\mathbf{b} - \mathbf{a}} \int_a^b f(t) dt \frac{1}{\mathbf{b} - \mathbf{a}} \int_a^b (\mathbf{t} - \mathbf{a})(\mathbf{b} - \mathbf{t}) dt \right| \\ \leq (\mathbf{b} - \mathbf{a}) \|\hat{f}\|_p \left| \frac{1}{(\mathbf{b} - \mathbf{a})} \int_a^b |(\mathbf{t} - \mathbf{a})(\mathbf{b} - \mathbf{t})|^2 dt - \left( \frac{1}{(\mathbf{b} - \mathbf{a})} \int_a^b |(\mathbf{t} - \mathbf{a})(\mathbf{b} - \mathbf{t})| dt \right)^2 \right|^{\frac{1}{p}} \end{aligned} \quad (8)$$

Since

$$\int_a^b |(\mathbf{t} - \mathbf{a})(\mathbf{b} - \mathbf{t})| dt = \frac{(\mathbf{b} - \mathbf{a})^3}{6}, \int_a^b f(t) dt = 1.$$

And

$$\int_a^b |(\mathbf{t} - \mathbf{a})(\mathbf{b} - \mathbf{t})|^2 dt = \int_a^b (\mathbf{t} - \mathbf{a})^2 (\mathbf{b} - \mathbf{t})^2 dt = \frac{(\mathbf{b} - \mathbf{a})^5}{30}.$$

Then (8) become,

$$\left| \frac{1}{b-a} \int_a^b f(t)(t-a)(b-t)dt - \frac{(b-a)^2}{6} \right| \leq (b-a) \|\hat{f}\|_p \left| \left( \frac{1}{(b-a)} \int_a^b \frac{(b-a)^5}{30} dt - \left( \frac{(b-a)^2}{6} \right)^p \right)^{\frac{1}{p}} \right|.$$

Then,

$$\left| \int_a^b f(t)(t-a)(b-t)dt - \frac{(b-a)^2}{6} \right| \leq (b-a)^2 \|\hat{f}\|_p \left[ \left( \frac{(b-a)^4}{30} - \left( \frac{(b-a)^2}{6} \right)^p \right)^{\frac{1}{p}} \right]$$

By using Lemma 1.2.1, we get

$$\left| |b - E(X)| |E(X) - a| - \sigma^2(X) - \frac{(b-a)^2}{6} \right| \leq (b-a)^2 \|\hat{f}\|_p \left[ \left( \frac{(b-a)^4}{30} - \left( \frac{(b-a)^2}{6} \right)^p \right)^{\frac{1}{p}} \right],$$

This implies,

$$\begin{aligned} & \left| |b - E(X)| |E(X) - a| - \sigma^2(X) - \frac{(b-a)^2}{6} \right| \leq \\ & (b-a)^2 \|\hat{f}\|_p 2^{\frac{1}{p}-1} \left( \left( \frac{(b-a)^4}{30} \right)^{\frac{1}{p}} + \frac{(b-a)^2}{6} \right) \\ & \leq (b-a)^2 \|\hat{f}\|_p 2^{\frac{1}{p}-1} \left( \frac{2^{\frac{4}{p}-1} \left( b^{\frac{4}{p}} + a^{\frac{4}{p}} \right)}{30^{\frac{1}{p}}} + \frac{2}{6} (b^4 + a^4) \right) \\ & \leq (b-a)^2 \|\hat{f}\|_p 2^{\frac{1}{p}-1} \left( b^{\frac{4}{p}} + a^{\frac{4}{p}} + b^4 + a^4 \right) \\ & \leq (b-a)^2 \|\hat{f}\|_p 2^{\frac{1}{p}-1} \left( b^{\frac{4}{p}} + a^{\frac{4}{p}} + b^{\frac{4}{p}} + a^{\frac{4}{p}} \right) \\ & \leq (b-a)^2 \|\hat{f}\|_p 2^{\frac{1}{p}} \left( b^{\frac{4}{p}} + a^{\frac{4}{p}} \right) \\ & \leq 2(b^2 + a^2) \|\hat{f}\|_p 2^{\frac{1}{p}} \left( b^{\frac{4}{p}} + a^{\frac{4}{p}} \right) \\ & \leq 2^{\frac{1}{p}+2} \left( b^{\frac{4}{p}} + a^{\frac{4}{p}} \right) \|\hat{f}\|_p \\ & \leq 2^{\frac{1}{p}+3} b^{\frac{4}{p}} \|\hat{f}\|_p \end{aligned}$$

Then ,we get

$$\left| |b - E(X)| |E(X) - a| - \sigma^2(X) - \frac{(b-a)^2}{6} \right| \leq 2^{\frac{1}{p}+3} b^{\frac{4}{p}} \|\hat{f}\|_p \quad \blacksquare$$

**Lemma 5.2.7[4]**

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Vol.7 No.3 (March, 2022)

If  $\mathbf{g}, \mathbf{h}, \dot{\mathbf{h}} \in L_1[a, b]$ , then

$$\left| \frac{1}{b-a} \int_a^b h(t)g(t)dt - \frac{1}{b-a} \int_a^b h(t)dt \frac{1}{b-a} \int_a^b g(t)dt \right|^2 \leq \left( \frac{b-a}{\pi} \right)^2 \int_a^b |\dot{h}|^2 \left[ \frac{1}{b-a} \int_a^b g(t)^2 dt - \left( \frac{1}{b-a} \int_a^b g(t)dt \right)^2 \right]$$

Is Pre-Lupas inequality, when  $\mathbf{g}, \mathbf{h}$  and  $\dot{\mathbf{h}} \in L_2[a, b]$

Now let us generalize Pre-Lupas inequality for  $L_p[a, b], p < 1$  spaces.

### Theorem 1.2.8

If  $\mathbf{g}, \mathbf{h}, \dot{\mathbf{h}} \in L_p[a, b], 0 < p < 1$ , then

$$\left| \frac{1}{b-a} \int_a^b h(t)g(t)dt - \frac{1}{b-a} \int_a^b h(t)dt \frac{1}{b-a} \int_a^b g(t)dt \right| \leq \frac{c(b-a)^{p-\frac{1}{p}}}{\pi} \|\dot{h}\|_p \left( \frac{1}{b-a} \int_a^b g(t)^2 dt - \int_a^b g(t)dt \right)$$

Where c is a positive constant.

### Proof

By using Lemma 1.2.7 and (2) we get,

$$\begin{aligned} \left| \frac{1}{b-a} \int_a^b h(t)g(t)dt - \frac{1}{b-a} \int_a^b h(t)dt \frac{1}{b-a} \int_a^b g(t)dt \right| &\leq \\ \frac{c(b-a)}{\pi} \left( \sum_{i=1}^n \frac{b-a}{n} |\dot{h}(t_i)|^2 \right)^{\frac{1}{2}} &\left( \frac{1}{b-a} \int_a^b g(t)^2 dt - \int_a^b g(t)dt \right) \end{aligned}$$

Then using (3) secondly, we get

$$\begin{aligned} \left| \frac{1}{b-a} \int_a^b h(t)g(t)dt - \frac{1}{b-a} \int_a^b h(t)dt \frac{1}{b-a} \int_a^b g(t)dt \right| &\leq \\ \frac{c(b-a)}{\pi} \|\dot{h}\|_p &\left( \frac{1}{b-a} \int_a^b g(t)^2 dt - \int_a^b g(t)dt \right) \quad \blacksquare \end{aligned}$$

### Collorally 1.2.9

$$\left| |b - E(X)| |E(X) - a| - \sigma^2(X) - \frac{(b-a)^3}{6} \right| \leq \frac{c(b-a)^{5+p-\frac{1}{p}}}{6\pi} \left( \frac{b-a}{5} - 1 \right) \|\dot{f}\|_p,$$

where c is a positive constant.

### Proof

Put  $\mathbf{h}(t) = \mathbf{f}(t), \mathbf{g}(x) = (\mathbf{t} - \mathbf{a})(\mathbf{b} - \mathbf{t})$  in Theorem 1.2.8, we get

$$\begin{aligned} \left| \frac{1}{b-a} \int_a^b f(t)(t-a)(b-t)dt - \frac{1}{b-a} \int_a^b f(t)dt \frac{1}{b-a} \int_a^b (t-a)(b-t)dt \right| &\leq \\ \frac{c(b-a)^{p-\frac{1}{p}}}{\pi} \|\dot{f}\|_p &\left( \frac{1}{b-a} \int_a^b [(t-a)(b-t)]^2 dt - \int_a^b (t-a)(b-t)dt \right) \quad (9) \end{aligned}$$

By using Lemma 1.2.1, we get

$$\int_a^b (\mathbf{b} - \mathbf{t}) (\mathbf{t} - \mathbf{a}) f(\mathbf{t}) d\mathbf{t} = |\mathbf{b} - E(X)| |E(X) - \mathbf{a}| - \sigma^2(X),$$

Since

$$\int_a^b (\mathbf{t} - \mathbf{a})^2 (\mathbf{b} - \mathbf{t})^2 d\mathbf{t} = \frac{(\mathbf{b} - \mathbf{a})^5}{30} \text{ and } \int_a^b f(\mathbf{t}) d\mathbf{t} = 1.$$

Also,

$$\int_a^b (\mathbf{b} - \mathbf{t}) (\mathbf{t} - \mathbf{a}) d\mathbf{t} = \frac{(\mathbf{b} - \mathbf{a})^3}{6},$$

So ,( 9) implies,

$$\left| |\mathbf{b} - E(X)| |E(X) - \mathbf{a}| - \sigma^2(X) - \frac{(\mathbf{b} - \mathbf{a})^3}{6} \right| \leq \frac{c(\mathbf{b} - \mathbf{a})^{2+p-\frac{1}{p}}}{\pi} \|\hat{f}\|_p \left( \frac{(\mathbf{b} - \mathbf{a})^4}{30} - \frac{(\mathbf{b} - \mathbf{a})^3}{6} \right).$$

Then , we get,

$$\left| |\mathbf{b} - E(X)| |E(X) - \mathbf{a}| - \sigma^2(X) - \frac{(\mathbf{b} - \mathbf{a})^3}{6} \right| \leq \frac{c(\mathbf{b} - \mathbf{a})^{5+p-\frac{1}{p}}}{6\pi} \left( \frac{\mathbf{b} - \mathbf{a}}{5} - 1 \right) \|\hat{f}\|_p,$$

where c is a positive constant ■

### Theorem 1.2.10

Let  $\mathbf{X}$  be a random variable with the probability density function  $f: [\mathbf{a}, \mathbf{b}] \rightarrow \mathbf{R}$ . If  $f \in L_p^n[\mathbf{a}, \mathbf{b}] = \{f: [\mathbf{a}, \mathbf{b}] \rightarrow \mathbf{R}, f, f^{(n)} \in L_p[\mathbf{a}, \mathbf{b}], 0 < p < 1\}$ , then we have

$$\left| |\mathbf{b} - E(X)| |E(X) - \mathbf{a}| - \sigma^2(X) - \sum_{i=0}^n \frac{(i+1)(\mathbf{b} - \mathbf{a})^{i+1} f^{(i)}(\mathbf{a})}{(i+3)!} \right| \leq \frac{c}{n!} \|f^{(n+1)}\|_p \frac{(\mathbf{b} - \mathbf{a})^{(n+1)p+3}}{(np+2)(np+3)},$$

where c is an absolute constant.

### Proof

The Taylors formula with integral remainder [13] is

$$f(t) = \sum_{i=0}^n \frac{(t - \mathbf{a})^n}{i!} f^{(i)}(\mathbf{a}) + \frac{1}{n!} \int_a^t (t - s)^n f^{(n+1)}(s) ds \\ t \in [\mathbf{a}, \mathbf{b}] \quad (10)$$

By using Lemma1.2.1 and (10), we have

$$|\mathbf{b} - E(X)| |E(X) - \mathbf{a}| - \sigma^2(X) \\ = \int_a^b (\mathbf{b} - \mathbf{t}) (\mathbf{t} - \mathbf{a}) \left[ \sum_{i=0}^n \frac{(\mathbf{t} - \mathbf{a})^n}{i!} f^{(i)}(\mathbf{a}) + \frac{1}{n!} \int_a^t (\mathbf{t} - s)^n f^{(n+1)}(s) ds \right] dt, \\ = \sum_{i=0}^n \frac{(\mathbf{t} - \mathbf{a})^n}{i!} f^{(i)}(\mathbf{a}) \int_a^b (\mathbf{b} - \mathbf{t}) (\mathbf{t} - \mathbf{a}) dt + \frac{1}{n!} \left[ \int_a^b (\mathbf{b} - \mathbf{t}) (\mathbf{t} - \mathbf{a}) \int_a^t (\mathbf{t} - s)^n f^{(n+1)}(s) ds \right] dt \quad (11)$$

Using the transformation  $t = (1 - u)\mathbf{a} + ub$ ,  $t \in [\mathbf{a}, \mathbf{b}]$

If  $t = \mathbf{a}$ , then

$$a = (1 - u)a + ub$$

$$a = a - au + ub$$

$$0 = u(b - a).$$

This implies,  $u = 0$  where  $t = \mathbf{a}$ .

Similarly,

If  $t = b$ , we obtain  $u = 1$ . Also,  $dt = (b - a)du$ ,

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Vol.7 No.3 (March, 2022)

$$\int_a^b (t-a)^{i+1} (b-t) dt = (b-a)^{i+3} \int_0^1 u^{i+1} (1-u) du = \frac{1}{(i+2)(i+3)}$$

By using (11), we deduce that,

$$|b - E(X)| |E(X) - a| - \sigma^2(X)$$

$$= \sum_{i=0}^n \frac{1}{(i+2)(i+3)} \frac{(b-a)^{i+1} f^{(i)}(a)}{i!} + \frac{1}{n!} \left[ \int_a^b (b-t) (t-a) \int_a^t (t-s)^n f^{(n+1)}(s) ds dt \right]$$

This implies,

$$\begin{aligned} & \left| |b - E(X)| |E(X) - a| - \sigma^2(X) - \sum_{i=0}^n \frac{(i+1)(b-a)^{i+1} f^{(i)}(a)}{(i+3)!} \right| \\ & \leq \frac{1}{n!} \int_a^b (b-t) (t-a) \left| \int_a^t (t-s)^n f^{(n+1)}(s) ds \right| dt \end{aligned} \quad (12)$$

Since

$$\left| \int_a^t (t-s)^n f^{(n+1)}(s) ds \right| \leq \int_a^t |t-s|^n |f^{(n+1)}(s)| ds.$$

Then by using (3), this implies,

$$\left| \int_a^t (t-s)^n f^{(n+1)}(s) ds \right| \leq c(b-a)^{p-\frac{1}{p}} \left( \int_a^t |t-s|^{pn} |f^{(n+1)}(s)|^p ds \right)^{\frac{1}{p}},$$

Then, we get,

$$\left| \int_a^t (t-s)^n f^{(n+1)}(s) ds \right| \leq c(b-a)^{p-\frac{1}{p}} |t-s|^{pn} \|f^{(n+1)}\|_p,$$

$$0 < p < 1. \quad (13)$$

Put (13) in (12), we get,

$$\begin{aligned} & \left| |b - E(X)| |E(X) - a| - \sigma^2(X) - \sum_{i=0}^n \frac{(i+1)(b-a)^{i+1} f^{(i)}(a)}{(i+3)!} \right| \\ & \leq \frac{c(b-a)^{p-\frac{1}{p}}}{n!} \|f^{(n+1)}\|_p \int_a^b (b-t) (t-a)^{pn+1} dt. \end{aligned} \quad (14)$$

Assume  $t = (1-u)a + ub$ .

So

$$\begin{aligned} & \left| |b - E(X)| |E(X) - a| - \sigma^2(X) - \sum_{i=0}^n \frac{(i+1)(b-a)^{i+1} f^{(i)}(a)}{(i+3)!} \right| \\ & \leq \frac{c(b-a)^{p-\frac{1}{p}}}{n!} \|f^{(n+1)}\|_p (b-a)^{np+3} \int_0^1 u^{np+1} (1-u) du \\ & \leq \frac{c(b-a)^{p-\frac{1}{p}}}{n!} \|f^{(n+1)}\|_p \frac{(b-a)^{np+3}}{(nP+2)(nP+3)} \\ & = \frac{c}{n!} \|f^{(n+1)}\|_p \frac{(b-a)^{(n+1)p+3-\frac{1}{p}}}{(nP+2)(nP+3)} \end{aligned}$$

This implies,

$$\begin{aligned} \left| |\mathbf{b} - E(X)| |E(X) - a| - \sigma^2(X) - \sum_{i=0}^n \frac{(i+1)(\mathbf{b}-a)^{i+1} f^{(i)}(a)}{(i+3)!} \right| &\leq \frac{c}{n!} \|f^{(n+1)}\|_p \frac{(\mathbf{b}-a)^{(n+1)p+3-\frac{1}{p}}}{(nP+2)(nP+3)} \\ &\leq \frac{c}{n!} \|f^{(n+1)}\|_p \frac{(\mathbf{b}-a)^{(n+1)p+3}}{(nP+2)(nP+3)} \end{aligned}$$

where c is absolute constant. ■

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